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# Analysis of variance in augmented block design and Scott-Knott's test in hybrid corn selection studies






**Abstract** – The objective of this work was to present a methodological alternative for studies of the characterization and selection of corn (*Zea mays*) genotypes, through the joint analysis of variance of an augmented block design, using Scott-Knott's test, and to present the hybrids selected from the adopted strategy, to show its efficiency. For the application of the methodology, a case study was used: the selection of superior corn hybrids for the Brazilian Cerrado. In four locations, 70 experimental hybrids were evaluated in an augmented block design without replicates, with three controls replicated once in each block. The analysis of experimental groups applied to the augmented block design, followed by genotype classifications by Scott-Knott's multiple comparison test, is a viable alternative for studies with a low number of replicates and a large number of genotypes. Based on the tested methodology, the following experimental hybrids are selected for grain yield: HT007, HT008, HT018, HT004, HT024, HT005, and HT071.

**Index terms:** *Zea mays*, grain yield, statistical analysis.

## Análise de variância em delineamento de blocos aumentados e teste de Scott-Knott em estudos de seleção de milho híbrido

**Resumo** – O objetivo deste trabalho foi apresentar uma alternativa metodológica para estudos de caracterização e seleção de genótipos de milho (*Zea mays*), por meio da análise de variância conjunta do delineamento de blocos aumentados, com uso do teste de Scott-Knott, e apresentar os híbridos selecionados a partir da estratégia adotada, para mostrar sua eficiência. Para aplicação da metodologia, utilizou-se um estudo de caso: a seleção de híbridos superiores de milho para o Cerrado. Em quatro locais, 70 híbridos experimentais foram avaliados em delineamento de blocos aumentados sem repetição, com três controles repetidos uma vez em cada bloco. A análise de grupos experimentais aplicada ao delineamento de blocos aumentados, seguida da classificação dos genótipos pelo teste de comparação múltipla de Scott-Knott, é uma alternativa viável para estudos com baixo número de repetições e grande quantidade de genótipos. Com base na metodologia testada, os seguintes híbridos experimentais são selecionados quanto ao rendimento de grãos: HT007, HT008, HT018, HT004, HT024, HT005 e HT071.

**Termos para indexação:** *Zea mays*, rendimento de grãos, análise estatística.

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## Introduction

The augmented block design proposed by Walter T. Federer, in 1955, sought to supply the needs of breeding programs in the preliminary experimentation phases, whose basic purpose is the screening of promising treatments for more accurate future tests (Duarte, 2000). A strong justification for this methodology indication is the little availability of propagation material for each new genotype to be evaluated, which forces the breeders not only to reduce the number of replicates, but also to adopt plots with only one or two rows of plants, without borders.

The analysis of variance method has its proposition attributed to Fischer (1919). Its classical version, historically the most widely applied, is employed for balanced data sets, in fixed or completely random models. This method produces estimators with desirable statistical properties, which are unbiased and with minimum variance (Barbin, 2019). The use of the augmented block design allows of two types of analysis, the intrablock analysis (fixed effects for blocks and treatments), and the analysis with recovery of interblock information (one with fixed effect and another with random effect) (Cochran & Cox, 1992).

The result analysis of an experimental group (called conjoint analysis) allows of more reliable conclusions and knowledge of the recommendation comprehensiveness (Pimentel Gomes & Guimarães, 1958).

To perform the pooled analysis of variance for the Federer's augmented block design, it is necessary that the experiments can be grouped without difficulty, for which it is necessary that the residual mean squares, obtained in the individual analysis of experiments, are not very different from each other, that is, that they have homogeneous variances for experimental errors (Pimentel-Gomes, 2000).

Scott-Knott's test is an alternative to a post-hoc analysis of variance, but has a different concept from the traditional multiple comparison tests such as those of Tukey, Duncan, SNK, and Dunnett. Scott-Knott's method separates the means of the treatments into homogeneous groups, thus minimizing the sum of squares within groups and maximizing it between them, without overlapping these groups. The objective of this method is to observe how important is the division of  $j$  treatments into homogeneous groups, based on the likelihood ratio, thus implying the

maximization of the sum of squares between the generated groups (Ramalho et al., 2000).

Regarding the generalized use of Scott-Knott's test, it should be understood that it is a method that was only proposed as an alternative to avoid ambiguity, in order to facilitate the interpretation of research results, especially in studies with a large number of treatments (Ferreira et al., 1999). However, it is important to clarify that ambiguity should not be understood as a problem in comparison tests because it is part of one of the possible outcomes predicted in such tests to detect differences.

Corn is of great importance among the cultivated grains, and Brazil is one of the its main producers in the world (Klein & Luna, 2022). The Midwest region of the country shows the highest rates of planted area and yield. In the 2020/2021 harvest, 21,584,400 ha were planted, which represents the second largest grain-planted area, and 114,691,300 Mg corn were harvested in Brazil (Conab, 2022).

The objective of this work was to present a methodological alternative for studies of the characterization and selection of corn genotypes, through the joint analysis of variance of an augmented block design, using Scott-Knott's test, and to present the hybrids selected from the adopted strategy, to show its efficiency.

## Materials and Methods

The use of the suggested methodology can be exemplified with data from an experiment installed in five sites of great representativeness, for the production of second-crop (safrinha) corn in Brazil used as a case study. Out of these five sites, two are located in a low-altitude region (<500 m), and three, in a high-altitude region (>500 m), and they are described as follows.

Site 1: located in the Fazenda Santo Antônio (12°52'S, 55°49'W, at 398 m altitude), in the municipality of Sorriso, in the state of Mato Grosso (MT), Brazil. Sowing date: 02/02/2018.

Site 2: located in Fazenda Santa Rita (15°28'S, 54°09'W, at 630 m altitude), in the municipality of Primavera do Leste, MT, Brazil. Sowing date: 01/30/2018.

Site 3: Fazenda Feijó (13°35'S, 52°21'W, at 344 m altitude), in the municipality of Canarana, MT, Brazil. Sowing date: 02/14/2018.

Site 4: Fazenda Casa Fértil (17°43'S, 50°58'W, at 825 m altitude), in the municipality of Rio Verde, in the state of Goiás (GO), Brazil. Sowing date: 02/20/2018.

Site 5: Fazenda Santa Bárbara (16°18'S, 47°48'W, at 1,020 m altitude), in the municipality of Cristalina, GO, Brazil. Sowing date: 01/25/2018.

In all locations, the experiments were installed in a rainfed (without irrigation), third-party area, and they were carried out together with the commercial corn area of the farm, after soybean was harvested. The experiment at site 5 was discarded because of phytotoxicity problems in one of the pesticide applications; and the other four sites remained for the evaluation of the experimental hybrids.

An augmented block delineation was carried out for the evaluation of 73 corn genotypes (triple hybrids), with 70 treatments without replicates (experimental hybrids), with 3 controls (20A55, MG699, and 2B633) replicated in each of the blocks, orthogonally arranged, so that each control was present in all rows and columns of the experiment (Figure 1). There were a total of 10 with 10 plots per block, totaling 100 plots.

The triple hybrids used as control are commercial hybrids, already known and of proven performance, which were acquired with own resources. The 70 new genotypes under test came from the breeding program of a company located in the municipality of Cristalina, GO, and they were obtained from a partial diallel, in which inbred lines were crossed with experimental simple hybrids, both from the germplasm bank of the company.

Seed of all treatments were treated with two insecticides, to ensure the initial pest protection for good stand formation of the plots. Two different classes of insecticides were used for seed treatment, one from the neonicotinoid class, and the other from the diamide class, at the doses recommended by the manufacturer, for the control of initial pests such as Hemiptera (bugs) and Lepidoptera (caterpillars), respectively.

The plot size was 4 rows at 0.5 m spacing by 5 m length (10 m<sup>2</sup>); the useful area was formed of the two central rows by 5 m length (5 m<sup>2</sup>). For the evaluation of corn for grain yield, Sturion et al. (1994) determined that plots of 3.64 m<sup>2</sup>, with approximately 20 plants, were sufficient to obtain a good precision in the evaluations.

The planting of the experiments was done manually with the aid of a planter, after the area had been mechanically furrowed by the planter. Approximately 15

days after plant emergence, thinning was performed to standardize the number of plants per plot. After thinning, the density was equivalent to 60,000 plants per hectare.

The fertilization and the cultural treatments followed the technical recommendations indicated for corn cultivation (Fancelli, 2015). Fungicides were not applied to any of the locations due to the low incidence of foliar diseases.

At harvesting, the grain weight per plot and the moisture content of the grain mass were attained. The data of weight per plot were converted into kilograms per hectare (yield) and corrected to 13% water content.

The statistical analysis of variance was used to evaluate the genotypes in the augmented block design, in which: the letter B refers to blocks; T refers to the control treatments (controls) and L refers to treatments (hybrids) (Table 1).

For this model, the total sum of squares ( $SS_{total}$ ) and block ( $SS_{blocks}$ , ignoring treatments) are calculated in the usual way, as follows:  $Y_{ij}$  are called the observations made in each plot;  $Y_j$  are the block totals, where  $j$  represents the number of blocks;  $i$  is the total number of treatments; and  $CF$  is the corrector factor. Thus, we have:

$$SS_{total} = \sum Y_{ij}^2 - CF$$

$$SS_{blocks} = \frac{1}{I} \sum Y_j^2 - CF$$

$$CF = \frac{1}{ij} \left( \sum Y_{ij} \right)^2$$

To calculate the sums of squares of adjusted treatments and their partitions, the model adjusted the values obtained for the hybrids under test. The resulting value from the difference between means of the controls of the block ( $MC_j$ ), where the treatment is located, and the general mean of the

**Table 1.** Statistical analysis of variance.

Source of variation	Degree of freedom (DF)
Block (ignoring treatments)	B-1
Treatment (adjusted)	T+L-1
Control	T-1
Hybrids	L-1
Control x Hybrids	1
Error	(B-1)(T-1)
Total	BT+L-1

B refers to blocks; T, control treatments; and L, treatments (hybrids).

controls (MC) in the whole experiment was added to the observed values of each hybrid. Thus, the sum of squares of the treatments (adjusted) is given by:

$$SS_{\text{treat.adj}} = \frac{1}{j} \sum T_i^2 + \sum L_k^2 - CF$$

Thus, an individual statistical analysis was performed for each of the sites, in order to determine the mean square error (MSE), and to know if it would be possible to perform the joint analysis of the experiments. With the results of the local MSEs, the homogeneity among the different locations was tested from the maximum F-test of Hartley, to proceed with the joint analysis (Hartley, 1950). To perform the joint analysis of variance for the Federer's augmented block design, these residual mean squares obtained in the individual analysis of experiments should not very different from each other, that is, that they should have homogeneous variances of experimental errors.

In the present study the intrablock analysis model was adopted. This model has the following assumptions: fixed effects for blocks and treatments, corresponding to the within-block analysis. The effects  $\mu$  (population average),  $\beta_j$  (index block j),  $\tau_i$  (index treatment i), and  $\varepsilon_{ij}$  (random error) were admitted as independent of each other, and the only variance component is associated with the experimental error, in accordance with the assumption  $\varepsilon_{ij} \sim N(0, \sigma_e^2)$ .

The statistical analyses were performed using the SAS Proc GLM software (SAS Institute Inc., 1997), where each of the four locations was analyzed individually. With the results of the MSE of each location, the joint analysis was determined. The output values were subjected to Scott-Knott's test built from a macro adapted for the SAS software, for each of the locations individually, and for the joint analysis.

## Results and Discussion

The results of the analysis of variance for each of the locations showed the following parameters: degrees of freedom (DF), mean squares (MS), and significance levels (F) data calculated for the sources of variation of the experiments (Table 2).

For the blocks source of variation (ignoring treatments), a significant difference was observed at 5% probability in Canarana (MT) and Sorriso (GO). In Primavera do Leste (MT) and Rio Verde (GO), the effect of blocks was not significant.

For the adjusted treatments, there was a significant difference at 5 and 1% probability in Primavera do Leste (MT) and Sorriso (GO), respectively. In Canarana (MT) and Rio Verde (GO), however, no significant difference was found for the treatments.

In all locations, there was a significant difference for the controls x hybrids interaction at 1% probability, except in Primavera do Leste (MT), where the interaction was not significant.

The MSEs values of the experiments were close, allowing of the joint analysis of the experiments. The obtained MSE values were: 700,980.74 in Canarana (MT), 851,792.87 in Primavera do Leste (MT), 1,412,143.67 in Rio Verde (GO), and 260,992.39 in Sorriso (GO). The maximum F was 5.25, calculated by dividing the highest MSE value by the lowest of the individual analyses. Since this value was below 7, it was possible to perform a joint analysis of the experiments (Hartley, 1950).

In the joint analysis of variance (Table 3), there was 1% probability for location, adjusted treatments, and for the controls x hybrids interaction. The significance level for location shows that there is a great difference from one location or region of evaluation to another. This confirms the good representativeness of the

**Table 2.** Analysis of F variance for the experimental sites in the municipalities located .

Source of variation	Degree of freedom	F (Canarana, MT)	F (Primavera do Leste, MT)	F (Rio Verde, GO)	F (Sorriso, GO)
Block (ignoring treatments)	9	2.61*	0.58 <sup>ns</sup>	2.17 <sup>ns</sup>	2.53*
Treatment (adjusted)	72	1.98 <sup>ns</sup>	2.55*	1.23 <sup>ns</sup>	5.36**
Control	2	1.55 <sup>ns</sup>	23.85**	0.49 <sup>ns</sup>	9.30**
Hybrids	69	1.20 <sup>ns</sup>	1.96 <sup>ns</sup>	0.55 <sup>ns</sup>	1.69 <sup>ns</sup>
Control x Hybrids	1	56.69**	0.43 <sup>ns</sup>	49.44**	250.88**
Error	18				
Total	171				

\* and \*\*Significant at 5% and 1% probability, respectively. <sup>ns</sup>Nonsignificant.



chosen locations, for the main producing regions of second-crop corn in the Brazilian Cerrado, which shows climatic and geographic differences and can be more or less affected by the rainfall regime in one or more locations, in one or more years, without having a tendency to do so. As to the treatments, significant differences between them were expected, which allows of the selection of the best genotypes in more locations and with more replicates for the following evaluation stages of the genetic improvement program.

There was a significant difference, at 5% probability, for blocks (ignoring treatments) and hybrids. For the controls, there was no significant difference, which was expected since the performance of this source of variation had already been observed in other genotypes-selection trials. This nonsignificance helps to increase the confidence in the conduction of the experiments and in the quality of the collected data. It also allows of the visualization of a good response of the statistical method used in this work, for the evaluation and selection of hybrids in the joint analysis for the augmented block design.

Another important observation for the results of the individual analysis of variances of the four trials are the obtained coefficients of variation (CVs).

According to Pimentel Gomes & Guimarães (1958), the CV gives an idea of the precision of the experiment. The highest and lowest CVs were 22.58% and 12.22% in Primavera do Leste (MT) and Sorriso (GO), respectively (Table 4).

When studying the CVs of several agricultural trials, Pimentel Gomes & Guimarães (1958) proposed the following CV classification: low, <10%; medium, between 10 and 20%; high, between 20 and 30%; and very high, > 30%. This classification is inversely proportional to that of the precision of the experiment,

i.e., the higher the CV the lower the experimental precision. However, since CVs can be misinterpreted (Doring & Reckling, 2018), their ranges of variation can vary and should be specific for each analyzed variable and for each studied crop (Fritsche-Neto et al., 2012; Couto et al., 2013).

Analyzing the CVs of 66 studies on corn, Scapim (1995) found values between 2.24 and 38.14% for the grain weight characteristic. The same author classified the CVs as: low, up to 10%; medium, from 10 to 22%; high, 22 to 28%; and, very high, above 28%. Therefore, there is a slight divergence between the CV ranking values proposed by Pimentel-Gomes and Scapim (Fritsche-Neto et al., 2012; Lopes et al., 2021).

In the joint analysis, the CV was 18.78%, which is considered the average for data of grain weight, proving the good conduction of the experiments in the field and, consequently, the reliability of the data obtained to proceed with the selection of hybrids.

The grain yield of the genotypes for each location and in the joint analysis are presented (Table 4). There are also the results of Scott-Knott's mean grouping test for all locations and the joint analysis. In the joint analysis, which considers all locations, seven experimental hybrids performed equal to the three controls, they are: HT007, HT008, HT018, HT004, HT024, HT005, HT071.

The performance of these hybrids in each of the four locations remained very stable, showing that they are performing equal or better than one or more controls, depending on the location evaluated. This way, it is possible to have a great confidence in the grouping results of the joint analysis.

Another fact conveying a high degree of reliability to the method of conjoint analysis is that the controls are in the group of higher grain yield. This fact was

**Table 3.** Analysis of variance for the joint analysis of the four municipalities.

Source of variation	Degree of freedom (DF)	Mean square (MS)	F
Local	3	208,266,337.60	201.39**
Block (ignoring treatments)	36	1,516,992.80	1.47*
Treatment (adjusted)	72	3,413,052.33	3.30**
Control	2	1,245,851.10	1.20 <sup>ns</sup>
Hybrids	69	1,527,708.30	1.48*
Control x Hybrids	1	137,836,191.10	133.28**
Error	288	1,034,150.00	
Total	471		

\* and \*\*Significant at 5% and 1% probability, respectively. <sup>ns</sup>Nonsignificant.

**Table 4.** Yield of the corn (*Zea mays*) hybrids genotypes and coefficients of variation in the joint analysis by municipality<sup>(1)</sup>.

Control	Hybrid	Grain yield (kg ha <sup>-1</sup> )				
		Joint Analysis	Canarana (MT)	Primavera do Leste (MT)	Sorriso (MT)	Rio Verde (GO)
Coefficient of variation (CV)	-	18.78%	12.57%	22.58%	12.22%	17.83%
1	20A55	6,421.22a	7,282.95b	5,740.40b	5,025.78b	7,635.77b
3	MG699	6,405.03a	7,653.47a	3,854.84c	6,022.26a	8,089.57a
10	HT007	6,384.17a	7,146.31b	6,088.40a	4,701.08c	7,600.90b
11	HT008	6,358.82a	7,040.66b	4,725.75b	5,049.03b	8,619.82a
21	HT018	6,166.41a	7,266.22b	6,228.65a	3,824.11e	7,346.69b
7	HT004	6,154.42a	6,993.99b	6,692.23a	3,596.57e	7,334.86b
26	HT024	6,137.88a	5,738.37e	6,738.01a	4,239.24d	7,835.90b
8	HT005	6,124.08a	6,873.42b	5,818.59b	4,539.43c	7,264.90b
2	2B633	6,107.77a	7,940.40a	2,946.35d	5,449.72a	8,094.62a
72	HT071	5,976.12a	6,362.78d	6,338.80a	5,212.74b	5,990.16e
9	HT006	5,863.10b	6,766.18c	6,079.71a	3,611.11e	6,995.38b
22	HT019	5,812.08b	7,242.96b	5,572.89b	4,579.20c	5,853.28e
12	HT009	5,811.90b	7,044.88b	5,378.08b	5,504.08a	5,320.54f
41	HT039	5,762.17b	8,061.43a	4,788.15b	3,763.89e	6,435.21d
57	HT056	5,740.74b	8,060.49a	4,063.48c	4,675.84c	6,163.13d
50	HT049	5,666.70b	7,541.02a	5,037.09b	3,299.58g	6,789.13c
45	HT044	5,660.24b	6,206.54d	6,202.61a	3,419.90f	6,811.93c
55	HT054	5,605.90b	7,122.86b	4,814.74b	3,441.51f	7,044.50b
29	HT027	5,602.01b	6,259.12d	6,671.26a	4,781.61c	4,696.06f
73	HT072	5,458.22c	7,950.00a	5,233.92b	3,430.16f	5,218.82f
47	HT046	5,442.30c	6,179.52d	4,988.41b	4,330.26d	6,271.00d
54	HT053	5,426.48c	5,528.86f	5,875.37b	4,063.84d	6,237.83d
52	HT051	5,390.92c	6,988.49b	4,927.96b	2,822.29h	6,824.94c
48	HT047	5,372.55c	6,304.79d	5,206.72b	2,783.88h	7,194.79b
60	HT059	5,371.14c	6,518.97c	4,811.76b	3,405.22f	6,748.62c
24	HT021	5,294.52c	5,962.83e	5,026.08b	3,568.53e	6,620.63c
58	HT057	5,277.96c	7,817.69a	2,692.80e	3,784.98e	6,816.37c
69	HT068	5,271.78c	6,743.39c	2,671.09e	3,570.29e	8,102.37a
25	HT023	5,252.38c	4,991.63f	5,308.69b	3,287.19g	7,422.01b
23	HT020	5,234.27c	6,404.51d	3,342.81d	4,485.20c	6,704.56cc
40	HT038	5,206.75c	6,967.00b	5,218.92b	3,044.82g	5,596.25e
42	HT040	5,176.22c	6,773.41c	2,686.19e	3,745.35e	7,499.92b
31	HT029	5,159.16c	6,734.47c	3,427.06d	3,723.81e	6,751.31c
71	HT070	5,159.16c	8,073.91a	2,609.39e	3,030.39g	6,922.97b
28	HT026	5,095.19d	7,219.28b	4,354.22c	3,443.12f	5,364.14
6	HT003	5,095.07d	6,101.57d	3,655.18c	3,818.33e	6,805.22c
66	HT065	5,064.93d	7,101.08b	3,823.31c	3,435.54f	5,899.81e
37	HT035	5,049.31d	5,027.35f	4,624.65b	3,711.99e	6,833.27c
46	HT045	5,026.64d	7,927.11a	3,556.23d	3,505.03f	5,118.20f
61	HT060	5,019.47d	5,836.55e	5,309.80b	3,237.73g	5,693.80e
44	HT043	4,970.82d	6,658.43c	3,300.63d	3,666.27e	6,257.96d
56	HT055	4,959.08d	6,430.97d	4,985.42b	3,399.35f	5,020.60f
67	HT066	4,953.33d	6,972.11b	3,207.72d	3,156.78g	6,476.72d
49	HT048	4,934.73d	6,931.11b	3,198.58d	2,490.27i	7,118.97b
38	HT036	4,928.44d	5,893.29e	2,272.06e	4,308.18d	7,240.22b
27	HT025	4,869.79d	6,532.37c	3,872.97c	4,452.79c	4,621.03f
13	HT010	4,850.19d	4,296.46g	2,703.71e	5,006.45b	7,394.14b
39	HT037	4,843.80d	6,167.49d	3,495.16d	4,238.60d	5,473.96f
18	HT015	4,797.64d	6,083.26d	5,690.17b	3,387.44f	4,029.71g
70	HT069	4,731.40e	6,187.25d	1,279.95e	5,207.65b	6,250.75d

Continuation...

**Table 4.** Continuation...

Control	Hybrid	Grain yield (kg ha <sup>-1</sup> )				
		Joint Analysis	Canarana (MT)	Primavera do Leste (MT)	Sorriso (MT)	Rio Verde (GO)
Coefficient of variation (CV)	-	18.78%	12.57%	22.58%	12.22%	17.83%
5	HT002	4,694.65e	4,778.03g	3,293.82d	4,992.51b	5,714.24e
16	HT013	4,649.63e	5,477.58f	3,966.38c	3,721.16e	5,433.39f
63	HT062	4,629.92e	6,634.78c	3,442.63d	3,445.72f	4,996.57f
17	HT014	4,580.91e	5,080.70f	3,940.22c	3,964.56d	5,338.17f
53	HT052	4,549.75e	7,713.45a	1,452.12e	3,740.11e	5,293.33f
30	HT028	4,499.89e	5,962.98e	2,954.63d	4,259.51d	4,822.44f
68	HT067	4,464.91e	5,175.84f	2,095.43e	4,185.42d	6,402.97d
19	HT016	4,440.97e	6,178.68d	3,716.19c	2,887.08h	4,981.93f
51	HT050	4,404.02f	5,527.49f	3,133.23d	3,097.54g	5,857.83e
43	HT042	4,389.50f	5,968.74e	2,996.28d	3,326.96g	5,266.01f
62	HT061	4,330.53f	5,499.38f	2,251.05e	3,653.48e	5,918.21e
59	HT058	4,316.90f	5,844.62e	2,056.69e	4,062.85d	5,303.45f
64	HT063	4,263.32f	4,832.13g	4,014.33c	2,657.92h	5,548.90e
20	HT017	4,250.12f	5,631.68e	3,321.66d	2,162.92i	5,884.21e
15	HT012	4,224.97f	4,668.59g	4,021.16c	2,907.74h	5,302.41f
32	HT030	4,216.54f	4,939.86f	4,008.23c	2,494.10i	5,423.96f
65	HT064	4,188.97f	5,693.40e	1,937.96e	3,219.54g	5,904.97e
34	HT032	4,143.39f	4,612.70g	3,445.23d	2,473.46i	6,042.15e
4	HT001	4,062.52f	5,394.84f	3,124.70d	2,729.05h	5,001.49f
36	HT034	3,979.43f	5,351.85f	2,252.74e	3,162.36g	5,150.76f
33	HT031	3,859.08f	3,673.72g	3,172.73d	4,547.56c	4,042.31g
14	HT011	3,847.68f	5,331.65f	2,199.73e	3,138.42g	4,720.90f
35	HT033	3,627.54f	4,503.80g	1,967.66e	2,853.35h	5,185.36f

<sup>(1)</sup>Means followed by equal letters in the columns belong to the same grouping, by Scott-Knott's test, at 5% probability.

already foreseen, and it is frequent in preliminary trials because the controls are genotypes that have already had their performance validated in several research trials, before entering the agricultural seed market. Regarding the genotypes evaluated in these preliminary trials, it is worth noting that they are not completely homogeneous because the processes that involve the purity of the lineages are not complete, or in advanced stages of homozygosity. Generally, in parallel with the preliminary trials, the same lineages that are parents of these genotypes are undergoing processes of characterization and purification seeking greater genotypic and phenotypic homogeneity, to obtain better homogeneity in hybrids that will be tested in network trials in subsequent years, if selected for intermediate and/or VCU trials. This homogeneity in the advanced hybrids brings an increment of grain yield because it improves the pattern and uniformity of plants that compose the evaluated plot, reducing the possible presence of atypical and/or dominated plants, which in practice reduces the grain yield that appear in the preliminary phases.

The experimental hybrids HT007, HT008, HT018, HT004, HT024, HT005, and HT071 are the most promising ones to present the best grain yields, showing equal performance in comparison to the controls, by Scott-Knott's test at 5% probability, through the joint analysis in augmented block design.

## Conclusions

1. The experimental corn (*Zea mays*) hybrids HT007, HT008, HT018, HT004, HT024, HT005, and HT071 show the same performance for yield by Scott-Knott's test, at 5% probability, and by the joint analysis of four locations in augmented block design.

2. The analysis of experimental groups applied with augmented block design, followed by the classification of the genotypes performed by Scott-Knott's multiple comparison test, is a viable alternative to supply the following needs: a low number of replicates and a large number of genotypes; Scott-Knott's test provides a more objective way for differentiating the best treatments in the experiments, thus, among the

multiple comparison tests normally described in the literature, Scott-Knott's test reduces the ambiguity in the grouping of the means, facilitating the analysis interpretation of the differences between treatments, during plant breeding program.

## References

- BARBIN, D. **Componentes de variância: teoria e aplicações**. 3.ed. Piracicaba: FEALQ, 2019.
- COCHRAN, W.G.; COX, G.M. **Experimental designs**. 2<sup>nd</sup> ed. New York: J. Wiley & Sons, 1992.
- CONAB. Companhia Nacional de Abastecimento. **Séries Históricas das safras**. Available at: <<https://www.conab.gov.br/info-agro/safras/serie-historica-das-safras?start=10>>. Accessed on: Sept. 4 2022.
- COUTO, M.F.; PETERNELLI, L.A.; BARBOSA, M.H.P. Classification of the coefficients of variation for sugarcane crops. **Ciência Rural**, v.43, p.957-961, 2013. DOI: <https://doi.org/10.1590/S0103-84782013000600003>.
- DORING, T.F.; RECKLING, M. Detecting global trends of cereal yield stability by adjusting the coefficient of variation. **European Journal of Agronomy**, v.99, p.30-36, 2018. DOI: <https://doi.org/10.1016/j.eja.2018.06.007>.
- DUARTE, J.B. **Sobre o emprego e a análise estatística do delineamento em blocos aumentados no melhoramento genético vegetal**. 2000. 293p. Tese (Doutorado) - Escola Superior de Agricultura Luiz de Queiroz, Universidade de São Paulo, Piracicaba. DOI: <https://doi.org/10.11606/T.11.2000.tde-28022002-100846>.
- FANCELLI, A.L. (ed.). **Milho: gestão e manejo**. Piracicaba: Ed. USP, 2015. 167p.
- FERREIRA, D.F.; MUNIZ, J.A.; AQUINO, L.H. de. Comparações múltiplas em experimentos com grande número de tratamentos - utilização do teste de Scott Knott. **Ciência e Agrotecnologia**, p.745-752, 1999.
- FISCHER, R.A. The correlation between relatives on the supposition of Mendelian inheritance. **Earth and Environmental Science Transactions of the Royal Society of Edinburgh**, v.52, p.399-433, 1919. DOI: <https://doi.org/10.1017/S0080456800012163>.
- FRITSCHÉ-NETO, R.; VIEIRA, R.A.; SCAPIM, C.A.; MIRANDA, G.V.; REZENDE, L.M. Updating the ranking of the coefficients of variation from maize experiments. **Acta Scientiarum. Agronomy**, v.34, p.99-101, 2012. DOI: <https://doi.org/10.4025/actasciagron.v34i1.13115>.
- HARTLEY, H.O. The maximum F-ratio as a short-cut test for heterogeneity of variance. **Biometrika**, v.37, p.308-312, 1950. DOI: <https://doi.org/10.2307/2332383>.
- KLEIN, H.S.; LUNA, F.V. The impact of the rise of modern maize production in Brazil and Argentina. **Historia Agraria**, v.86, p.1-38, 2022. DOI: <https://doi.org/10.26882/histagrar.086e09k>.
- LOPES, B.G.; FARIA, G.A.; MALTONI, K.L.; ROCHA, P.S.; PEIXOTO, A.P.B.; OLIVEIRA, T.A. de; FONSECA, A.D. da; FELIZARDO, L.M. Classification of the coefficient of variation for experiments with eucalyptus seedlings in greenhouse. **Revista Ciência Agronômica**, v.52, e20207587, 2021. DOI: <https://doi.org/10.5935/1806-6690.20210050>.
- PIMENTEL GOMES, F.; GUIMARÃES, R.F. Joint analysis of experiments in complete randomised blocks with some common treatments. **Biometrics**, v.14, p.521-526, 1958. DOI: <https://doi.org/10.2307/2527518>.
- PIMENTEL-GOMES, F. **Curso de estatística experimental**. 14.ed. Piracicaba: Esalq/USP, 2000.
- RAMALHO, M.A.P.; FERREIRA, D.F.; OLIVEIRA, A.C. de. **Experimentação em genética e melhoramento de plantas**. Lavras: UFLA, 2000. p.87-134.
- SAS Institute. **SAS/STAT software: changes and enhancements through release 6.12**. Cary, 1997.
- SCAPIM, C.A.; CARVALHO, C.G.P. de; CRUZ, C.D. Uma proposta de classificação dos coeficientes de variação para a cultura do milho. **Pesquisa Agropecuária Brasileira**, v.30, p.683-686, 1995.
- STURION, J.A.; RESENDE, M.D.V. de; CARPANEZZI, A.A.; ZANON, A. Variação genética e seleção para características de crescimento em teste de progênes de *Mimosa scabrella* VAR. *aspericarpa*. **Boletim de Pesquisa Florestal**, n.28/29, p.73-83, 1994.